

Econ 220B

Answers to 220B Winter 2008 Final Exam

1a.) $(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$

b.) $(\mathbf{Rb})' \left[\mathbf{R}s^2(\mathbf{X}'\mathbf{X})^{-1}\mathbf{R}' \right]^{-1} (\mathbf{Rb})$

$$\mathbf{R} = \begin{bmatrix} 1 & -2 \\ & \end{bmatrix}$$

$F(1, T - 2)$

c.)

$$\frac{\mathbf{Rb}}{\sqrt{\mathbf{R}s^2(\mathbf{X}'\mathbf{X})^{-1}\mathbf{R}'}} \sim t(T - k)$$

d.) Replace $s^2(\mathbf{X}'\mathbf{X})^{-1}$ with

$$(\mathbf{X}'\mathbf{X})^{-1} \sum_{t=1}^T e_t^2 \mathbf{x}_t \mathbf{x}_t' (\mathbf{X}'\mathbf{X})^{-1}$$

e.) Let $g(\mathbf{b}) = \log b_1 - 2 \log b_2$

$$\mathbf{g}' = \begin{bmatrix} \frac{1}{b_1} & \frac{-2}{b_2} \end{bmatrix}$$

$$g(\mathbf{b})^2 / [\mathbf{g}' s^2(\mathbf{X}'\mathbf{X})^{-1} \mathbf{g}] \approx \chi^1(1)$$

2a.) Factors such as education might be correlated with both ε_i and s_i

b.) $E(\mathbf{x}_i [1 \ s_i])$ has rank 2. To test, regress s_i on \mathbf{x}_i and test whether coefficients on \mathbf{x}_i are all the same

c.) $E(\mathbf{x}_i \varepsilon_i) = \mathbf{0}$, could test with GMM test of overidentifying restrictions

d.) Step 1: regress s_i on \mathbf{x}_i , save fitted values \hat{s}_i

Step 2: regress y_i on constant and \hat{s}_i

3a.)

$$\frac{\partial \mathcal{L}}{\partial p} = \frac{N}{p} - \frac{T - N}{1 - p} = 0 \text{ or } \hat{p} = N/T$$

b.) $\sqrt{T}(\hat{p} - p) \xrightarrow{L} N(0, p(1 - p))$

c.)

$$h_t = \frac{y_t}{p} - \frac{1 - y_t}{1 - p} = \frac{y_t - p}{p(1 - p)}$$

d.)

$$E(h_t) = \frac{E(y_t) - p}{p(1 - p)} = \frac{p - p}{p(1 - p)} = 0$$

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e.)

$$S = E(h_t^2) = E \left[\frac{y_t^2 - 2py_t + p^2}{p^2(1-p)^2} \right] = \left[\frac{p - 2p^2 + p^2}{p^2(1-p)^2} \right] = \frac{1}{p(1-p)}$$

f.)

$$\begin{aligned} \frac{\partial h_t}{\partial p} &= \frac{-y_t}{p^2} - \frac{(1-y_t)}{(1-p)^2} \\ \frac{\partial g}{\partial p} &= -T^{-1} \sum_{t=1}^T \left[\frac{y_t}{p^2} + \frac{(1-y_t)}{(1-p)^2} \right] \\ \frac{\partial g}{\partial p} \Big|_{p=\hat{p}} &= - \left[\frac{1}{\hat{p}} + \frac{1}{1-\hat{p}} \right] = - \frac{1}{\hat{p}(1-\hat{p})} \end{aligned}$$

g.)

$$\hat{V} = \{ [\hat{p}(1-\hat{p})]^{-1} [\hat{p}(1-\hat{p})] [\hat{p}(1-\hat{p})]^{-1} \}^{-1} = \hat{p}(1-\hat{p})$$

which is same answer as for (b)

h.) Yes because the moment condition $E(h_t) = 0$ would still hold

i.) For b, plim is same but distribution is now $\sqrt{T}(\hat{p} - p) \xrightarrow{L} N(0, Q)$ for

$$Q = \sum_{j=-\infty}^{\infty} E(y_t - p)(y_{t-j} - p)$$

For (e) we have

$$\begin{aligned} S &= \sum_{j=-\infty}^{\infty} E \left[\frac{y_t - p}{p(1-p)} \right] \left[\frac{y_{t-j} - p}{p(1-p)} \right] \\ &= \left[\frac{1}{p^2(1-p)^2} \right] Q \end{aligned}$$

and

$$V = \left\{ [p(1-p)]^{-1} \left[\frac{Q}{p^2(1-p)^2} \right]^{-1} [p(1-p)]^{-1} \right\}^{-1} = Q$$

again the same answer.